

國立交通大學

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碩士論文

類別製造過程的多變量指數加權移動平均監控計劃

Multivariate Exponentially Weighted Moving Average
Monitoring Scheme for a Categorical Manufacturing Process



研究生：洪鏡婷

指導教授：陳志榮 博士

中華民國九十九年六月

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研究生：洪鏡婷

Student : Jing-Ting Hong

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Advisor : Dr. Chih-Rung Chen



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本文是發展一個方法去監控不同類別的資料。首先介紹 Dirichlet-多項式模型，根據該模型提出了類別製造過程的多變量指數加權移動平均管制圖並研究其相關的特性。最後用模擬的資料來說明這個方法的實用性與適用性。

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Student: Jing-Ting Hong

Advisor: Dr. Chih-Rung Chen

Institute of Statistics,
National Chiao Tung University



ABSTRACT

This paper is to develop a method to monitor the fractions of the tested items falling into different categories of pass/fail modes. The Dirichlet-compound multinomial model is first introduced and then a multivariate exponentially weighted moving average control chart is proposed for a categorical manufacturing process under the proposed model. Some relevant properties of the proposed control chart are also investigated. Finally, a simulation study is presented to show the usefulness and applicability of the proposed methodology.

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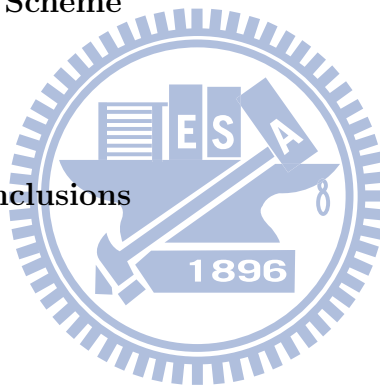
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1 Introduction

1.1 Motivation

In the current international marketplace, quality improvement is a key point for maintaining competitive advantage. Statistical process control is an effective tool for achieving process stability and improving process capability through variation reduction. When a product item is tested, usually one has more information than just pass or fail. Often there are categories of failures. For instance, a product may have several categories of failure modes. Thus, in the paper, a multivariate exponentially weighted moving average monitoring scheme is proposed for a categorical manufacturing process under the Dirichlet-compound multinomial model and then some relevant properties of the proposed monitoring scheme are also investigated.

1.2 Literature Review

Statistical process control (SPC) refers to some statistical methods which are widely used to monitor and improve the quality and productivity of industrial processes and service operations. SPC primarily involves the implementation of control charts. The method of control charts is a graphical tool which is used to monitor processes in order to distinguish special significant causes of variation from general assignable causes of variation in processes. The Shewhart (Shewhart, 1931), cumulative sums (CUSUM) (Page, 1954), and exponentially weighted moving average (EWMA) (Roberts, 1959) control charts are widely used in practice. Standard control chart usage involves phase I and phase II applications, with two different and distinct objectives. In phase I, a set of process data is gathered and analyzed all at once in a retrospective analysis, constructing trial control limits to determine if the process has been in control over the period of time where the data were collected, and to see if reliable control limits can be established to monitor future production. In phase II, we use the control chart to monitor the process by comparing the sample statistic for each successive sample as it is drawn from the process to the control limits.

The Shewhart control chart (Shewhart, 1931) monitors the process observations directly. Suppose that $\{x_{tm}: t = 1, 2, \dots \text{ and } m = 1, \dots, n\}$ are independent univariate measurements, where n is a known positive integer and x_{t1}, \dots, x_{tn} are identically distributed for each t . Set $\bar{x}_t \equiv \sum_{m=1}^n x_{tm}/n$ for $t = 1, 2, \dots$. Let μ_0 denote the known real-valued in-control process target and σ_0 the known positive in-control process standard deviation. When $\bar{x}_t \sim N(\mu_0, \sigma_0^2/n)$, the process is called in control at time t ; otherwise, out of control at time t . Then the Shewhart control chart is based on the values

$$w_t = \bar{x}_t - \mu_0 \quad (1)$$

for $t = 1, 2, \dots$. Shewhart (1931) proposed the stopping time of the Shewhart control chart as the first time t such that

$$|w_t| > L \frac{\sigma_0}{\sqrt{n}}, \quad (2)$$

where $L = 3$. In practice, L is often chosen to achieve a specified in-control average run length.

Page (1954) introduces the CUSUM chart as a sequential probability test. Suppose that $\{x_{tm}: t = 1, 2, \dots \text{ and } m = 1, \dots, n\}$ are independent univariate measurements, where n is a known positive integer and x_{t1}, \dots, x_{tn} are identically distributed for each t . Set $\bar{x}_t \equiv \sum_{m=1}^n x_{tm}/n$ for $t = 1, 2, \dots$. Let μ_0 denote the known real-valued in-control process target and σ_0 the known positive in-control process standard deviation. When $\bar{x}_t \sim N(\mu_0, \sigma_0^2/n)$, the process is called in control at time t ; otherwise, out of control at time t . The CUSUM algorithm assigns equal weights to past observations, and its tabular form consists of two quantities,

$$w_t^+ = \max [0, w_{t-1}^+ + (\bar{x}_t - \mu_0) - k\sigma_0/\sqrt{n}] \quad (3)$$

and

$$w_t^- = \min [0, w_{t-1}^- + (\bar{x}_t - \mu_0) + k\sigma_0/\sqrt{n}] \quad (4)$$

for $t = 1, 2, \dots$, where $w_0^+ = w_0^- = 0$ and k is the reference value which is often chosen about halfway between the target μ_0 and the out-of-control mean value μ_1 of interest.

Page (1954) proposed the stopping time of the CUSUM control chart as the first time t such that either w_t^+ or w_t^- exceed the decision interval H , where H is chosen to achieve a specified in-control average run length, e.g., $H = 5$ when $ARL_0 = 465$.

Roberts (1959) proposed a monitoring scheme which is based on the EWMA of the observations. The EWMA, originally called geometric moving average (GMA) in Roberts (1959), is briefly introduced as follows: Suppose that $\{x_{tm}: t = 1, 2, \dots \text{ and } m = 1, \dots, n\}$ are independent univariate measurements, where n is a known positive integer and x_{t1}, \dots, x_{tn} are identically distributed for each t . Set $\bar{x}_t \equiv \sum_{m=1}^n x_{tm}/n$ for $t = 1, 2, \dots$. Let μ_0 denote the known in-control real-valued process target and σ_0 the known positive in-control process standard deviation. When $\bar{x}_t \sim N(\mu_0, \sigma_0^2/n)$, the process is called in control at time t ; otherwise, out of control at time t . Then the EWMA control chart is based on the values

$$\begin{aligned}
 w_t &\equiv (1 - \lambda)w_{t-1} + \lambda(\bar{x}_t - \mu_0) \\
 &= \lambda(1 - \lambda)^{t-1}(\bar{x}_1 - \mu_0) + \dots + \lambda(1 - \lambda)(\bar{x}_{t-1} - \mu_0) + \lambda(\bar{x}_t - \mu_0) \\
 &= \lambda \sum_{i=0}^{t-1} (1 - \lambda)^i (\bar{x}_{t-i} - \mu_0)
 \end{aligned} \tag{5}$$

for $t = 1, 2, \dots$, where $w_0 \equiv 0$ and λ is a specified value in $(0, 1]$. If $\bar{x}_t \sim N(\mu_0, \sigma_0^2/n)$, then $w_t \sim N(0, \sigma_t^2)$, where

$$\sigma_t = \sqrt{\frac{\lambda [1 - (1 - \lambda)^{2t}]}{n(2 - \lambda)}} \sigma_0 \rightarrow \sqrt{\frac{\lambda}{n(2 - \lambda)}} \sigma_0 \tag{6}$$

as $t \rightarrow \infty$. Roberts (1959) proposed the stopping time of the EWMA monitoring scheme as the first time t such that

$$|w_t| > L\sigma_t, \tag{7}$$

where $L = 3$. In practice, L is often chosen to achieve a specified in control average run length. It is the same as the Shewhart control chart when $\lambda = 1$, and nearly the same as the CUSUM control chart when $\lambda \rightarrow 0$. There have been numerous extensions and variations of the basic EWMA control chart.

Through modern technology that allows simultaneously monitoring all key quality characteristics during a manufacturing process, the monitored quality characteristics are

usually dependent each other. This is especially true for quality characteristics related to safety, fault detection and diagnosis, quality control, and process control. Joint monitoring of quality characteristics ensures appropriate control of the overall process. Multivariate SPC techniques have recently been applied to novel fields such as environmental monitoring and detection of computer intrusion. The purpose of multivariate on-line techniques is to investigate whether quality characteristics are simultaneously in control or not. Versions of the multivariate Shewhart, CUSUM, and EWMA control charts have been proposed under the multivariate normality assumption.

To incorporate the recent historical information, Lowry *et al.* (1992) proposed a multivariate exponentially weighted moving average (MEWMA) control chart which is briefly introduced as follows: Suppose that $\{x_{tm}: t = 1, 2, \dots \text{ and } m = 1, \dots, n\}$ are independent p -variate measurements, where n is a known positive integer; x_{t1}, \dots, x_{tn} are identically distributed for each t ; and p (≥ 2) is a known positive integer. Set $\bar{x}_t \equiv \sum_{m=1}^n x_{tm}/n$ for $t = 1, 2, \dots$. Let μ_0 denote the known $p \times 1$ in-control process target vector in $(-\infty, \infty)^p$ and Σ_0 the known $p \times p$ positive definite in-control process covariance matrix. Let $N_p(\mu, \Sigma)$ denote the p -variate normal distribution with mean vector μ and covariance matrix Σ . When $\bar{x}_t \sim N_p(\mu_0, \Sigma_0/n)$, the process is called in control at time t ; otherwise, out of control at time t . Lowry *et al.* (1992) proposed the MEWMA control chart as based on the $p \times 1$ vectors

$$\begin{aligned}
 w_t &\equiv (I_p - \Lambda)w_{t-1} + \Lambda(\bar{x}_t - \mu_0) \\
 &= \Lambda(I_p - \Lambda)^{t-1}(\bar{x}_1 - \mu_0) + \dots + \Lambda(I_p - \Lambda)(\bar{x}_{t-1} - \mu_0) + \Lambda(\bar{x}_t - \mu_0) \\
 &= \Lambda \sum_{i=0}^{t-1} (I_p - \Lambda)^i (\bar{x}_{t-i} - \mu_0)
 \end{aligned} \tag{8}$$

for $t = 1, 2, \dots$, where $w_0 \equiv 0_{p \times 1}$, the $p \times 1$ vector $(0, \dots, 0)^T$; I_p denotes the identity matrix of order p ; and Λ is a specified diagonal matrix $\text{diag}\{\lambda_1, \dots, \lambda_p\}$ with $\lambda_1, \dots, \lambda_p \in (0, 1]$. Set $\Sigma_0 \equiv (\Sigma_{0jj'})_{j,j'=1,\dots,p}$. If $\bar{x}_t \sim N_p(\mu_0, \Sigma_0/n)$, then $w_t \sim N_p(0_{p \times 1}, \Sigma_t)$, where

$$\begin{aligned}
 \Sigma_t &\equiv (\Sigma_{tjj'})_{j,j'=1,\dots,p} = \frac{1}{n} \Lambda \left[\sum_{i=0}^{t-1} (I_p - \Lambda)^i \Sigma_0 (I_p - \Lambda)^i \right] \Lambda \\
 &\rightarrow \frac{1}{n} \Lambda \left[\sum_{i=0}^{\infty} (I_p - \Lambda)^i \Sigma_0 (I_p - \Lambda)^i \right] \Lambda
 \end{aligned} \tag{9}$$

as $t \rightarrow \infty$ with

$$\Sigma_{tjj'} = \frac{\lambda_j \lambda_{j'} [1 - (1 - \lambda_j)^t (1 - \lambda_{j'})^t] \Sigma_{0jj'}}{n(\lambda_j + \lambda_{j'} - \lambda_j \lambda_{j'})} \rightarrow \frac{\lambda_j \lambda_{j'} \Sigma_{0jj'}}{n(\lambda_j + \lambda_{j'} - \lambda_j \lambda_{j'})} \quad (10)$$

as $t \rightarrow \infty$. In particular, when $\lambda_1 = \dots = \lambda_p = \lambda$,

$$\Sigma_t = \frac{\lambda [1 - (1 - \lambda)^{2t}]}{n(2 - \lambda)} \Sigma_0 \rightarrow \frac{\lambda}{n(2 - \lambda)} \Sigma_0 \quad (11)$$

as $t \rightarrow \infty$. Then the stopping time of the MEWMA monitoring scheme is the first time t such that

$$w_t^T \Sigma_t^{-1} w_t (\equiv T_t^2) > h, \quad (12)$$

where h is chosen to achieve a specified in-control average run length.

Consider a manufacturing process where each product units can be classified as one of $k + 1$ disjoint categories for some fixed $k \in \{1, 2, \dots\}$. When the outcome is recorded as one of two categories, e.g., {pass, fail}, the data are called binary. When the outcome is recorded as one of $k + 1$ disjoint categories for some $k \in \{2, 3, \dots\}$, the data are called polytomous, e.g., {pass, the first defect type, ..., the k th defect type}. See, e.g., McCullagh and Nelder (1989). Several researchers have investigated categorical data in different situations. Shiau *et al.* (2005) proposed the Dirichlet-compound multinomial empirical Bayes model to monitor the polytomous data. In the paper, we develop an MEWMA control chart for monitoring a manufacturing process under the Dirichlet-compound multinomial model.

1.3 Outline

The paper is organized as follows. In Section 2, the Dirichlet-compound multinomial model for a categorical manufacturing process is briefly introduced. In Section 3, a multivariate exponentially weighted moving average control chart is proposed and then some relevant properties of the proposed control chart are also investigated. In Section 4, a simulation study is presented to illustrate the proposed methodology. Finally, comparison and conclusions are given in Section 5.

2 Dirichlet-Compound Multinomial Model for a Categorical Manufacturing Process

Consider a manufacturing process which produces product units having k different types of defects for some known positive integer k . In a product unit, let p_{it} denote the probability of having the i th defect type at time t for $i = 1, \dots, k$. Then $1 - \sum_{i=1}^k p_{it}$ ($\equiv p_{0t}$) is the probability of having none of these k defect types at time t . For $i = 1, \dots, k$, let x_{it} denote the number of tested product units having the i th defect type among n_t randomly chosen tested product units at time t . Then $n_t - \sum_{i=1}^k x_{it}$ ($\equiv x_{0t}$) is the number of tested product units having none of these k defect types at time t . Set $p_t \equiv (p_{0t}, p_{1t}, \dots, p_{kt})^T$ and $x_t \equiv (x_{0t}, x_{1t}, \dots, x_{kt})^T$. Then $p_t \in \mathcal{P}$ and $x_t \in \mathcal{X}_t$, where $\mathcal{P} \equiv \{p_t: p_{0t}, p_{1t}, \dots, p_{kt} \in (0, 1) \text{ with } \sum_{i=0}^k p_{it} = 1\}$ and $\mathcal{X}_t \equiv \{x_t: x_{0t}, x_{1t}, \dots, x_{kt} \in \{0, 1, \dots, n_t\} \text{ with } \sum_{i=0}^k x_{it} = n_t\}$. Then the number of elements in \mathcal{X}_t is $(n_t + k)! / (n_t! k!) \equiv |\mathcal{X}_t|$. Assume that x_t given p_t is distributed as either binomial($n_t; p_t$) for $k = 1$ or multinomial($n_t; p_t$) for $k \geq 2$. Then the conditional probability mass function (p.m.f.) of x_t given p_t is

$$f(x_t | p_t) = \frac{n_t!}{x_{0t}! x_{1t}! \cdots x_{kt}!} p_{0t}^{x_{0t}} p_{1t}^{x_{1t}} \cdots p_{kt}^{x_{kt}} \cdot 1_{\mathcal{X}_t}(x_t), \quad (13)$$

where $1_{\mathcal{X}_t}(x_t) = 1$ for $x_t \in \mathcal{X}_t$ and 0 otherwise.

Suppose that p_t is distributed as either beta(α) for $k = 1$ or Dirichlet(α) for $k \geq 2$, where $\alpha \equiv (\alpha_0, \alpha_1, \dots, \alpha_k)^T$ is the unknown $(k+1) \times 1$ parameter vector in the parameter space $(0, \infty)^{k+1}$. Set $\alpha_s \equiv \sum_{i=0}^k \alpha_i$. Then the probability density function (p.d.f.) of p_t is

$$f(p_t; \alpha) = \frac{\Gamma(\alpha_s)}{\Gamma(\alpha_0) \Gamma(\alpha_1) \cdots \Gamma(\alpha_k)} p_{0t}^{\alpha_0-1} p_{1t}^{\alpha_1-1} \cdots p_{kt}^{\alpha_k-1} \cdot 1_{\mathcal{P}}(p_t), \quad (14)$$

where $1_{\mathcal{P}}(p_t) = 1$ for $p_t \in \mathcal{P}$ and 0 otherwise. Then p_t given x_t is distributed as either beta($\alpha_0 + x_{0t}, \alpha_1 + x_{1t}$) for $k = 1$ or Dirichlet($\alpha_0 + x_{0t}, \alpha_1 + x_{1t}, \dots, \alpha_k + x_{kt}$) for $k \geq 2$ with

$$E_{\alpha}(p_t | x_t) = \left(\frac{\alpha_0 + x_{0t}}{\alpha_s + n_t}, \frac{\alpha_1 + x_{1t}}{\alpha_s + n_t}, \dots, \frac{\alpha_k + x_{kt}}{\alpha_s + n_t} \right)^T. \quad (15)$$

See, e.g., p. 217 of Johnson *et al.* (1995) for the parametric family of beta distributions and p. 488 of Kotz *et al.* (2000) for the parametric family of Dirichlet distributions.

Then the p.m.f. of x_t is

$$\begin{aligned} f(x_t; \alpha) &= \frac{f(x_t, p_t; \alpha)}{f(p_t|x_t; \alpha)} = \frac{f(x_t|p_t)f(p_t; \alpha)}{f(p_t|x_t; \alpha)} \\ &= \exp \left[\sum_{j=0}^{n_t-1} \log \left(\frac{j+1}{\alpha_s + j} \right) - \sum_{i=0}^k \sum_{j=0}^{x_{it}-1} \log \left(\frac{j+1}{\alpha_i + j} \right) \right] \cdot 1_{\mathcal{X}_t}(x_t); \end{aligned} \quad (16)$$

see, e.g., pp. 80–81 of Johnson *et al.* (1997). Then, given x_t , the likelihood function for α is

$$L(\alpha; x_t) = \exp \left[\sum_{j=0}^{n_t-1} \log \left(\frac{j+1}{\alpha_s + j} \right) - \sum_{i=0}^k \sum_{j=0}^{x_{it}-1} \log \left(\frac{j+1}{\alpha_i + j} \right) \right], \quad (17)$$

the log-likelihood function for α is

$$l(\alpha; x_t) = \log [L(\alpha; x_t)] = \sum_{j=0}^{n_t-1} \log \left(\frac{j+1}{\alpha_s + j} \right) - \sum_{i=0}^k \sum_{j=0}^{x_{it}-1} \log \left(\frac{j+1}{\alpha_i + j} \right), \quad (18)$$

the score function for α is

$$\begin{aligned} &\frac{\partial l(\alpha; x_t)}{\partial \alpha} \\ &= \left(\sum_{j=0}^{x_{0t}-1} \frac{1}{\alpha_0 + j}, \sum_{j=0}^{x_{1t}-1} \frac{1}{\alpha_1 + j}, \dots, \sum_{j=0}^{x_{kt}-1} \frac{1}{\alpha_k + j} \right)^T - \left(\sum_{j=0}^{n_t-1} \frac{1}{\alpha_s + j} \right) \cdot 1_{(k+1) \times 1} \\ &\equiv S(\alpha; x_t) \equiv (S_0(\alpha; x_t), S_1(\alpha; x_t), \dots, S_k(\alpha; x_t))^T, \end{aligned} \quad (19)$$

the observed Fisher information for α is

$$\begin{aligned} -\frac{\partial^2 l(\alpha; x_t)}{\partial \alpha \partial \alpha^T} &= \text{diag} \left\{ \sum_{j=0}^{x_{0t}-1} \frac{1}{(\alpha_0 + j)^2}, \sum_{j=0}^{x_{1t}-1} \frac{1}{(\alpha_1 + j)^2}, \dots, \sum_{j=0}^{x_{kt}-1} \frac{1}{(\alpha_k + j)^2} \right\} \\ &\quad - \left[\sum_{j=0}^{n_t-1} \frac{1}{(\alpha_s + j)^2} \right] \cdot 1_{(k+1) \times 1} 1_{(k+1) \times 1}^T \\ &\equiv J(\alpha; x_t) \equiv (J_{ii'}(\alpha; x_t))_{i,i'=0,1,\dots,k}, \end{aligned} \quad (20)$$

and the expected Fisher information for α is

$$Cov_{\alpha}(S(\alpha; x_t)) \equiv I_t(\alpha) \equiv (I_{ii'}(\alpha))_{i,i'=0,1,\dots,k}, \quad (21)$$

where $1_{(k+1) \times 1}$ denotes the $(k+1) \times 1$ vector $(1, \dots, 1)^T$. Notice that $E_{\alpha}(S(\alpha; x_t)) = 0_{(k+1) \times 1}$ and that $Cov_{\alpha}(S(\alpha; x_t)) = E_{\alpha}(J(\alpha; x_t))$, where $0_{(k+1) \times 1}$ denotes the $(k+1) \times 1$ vector

$(0, \dots, 0)^T$. Then

$$\begin{aligned} I_{tii}(\alpha) &= \sum_{x_t \in \mathcal{X}_t} \left(\sum_{j=0}^{x_{it}-1} \frac{1}{\alpha_i + j} - \sum_{j=0}^{n_t-1} \frac{1}{\alpha_s + j} \right)^2 f(x_t; \alpha) \\ &= \sum_{x_t \in \mathcal{X}_t} \left[\sum_{j=0}^{x_{it}-1} \frac{1}{(\alpha_i + j)^2} \right] f(x_t; \alpha) - \sum_{j=0}^{n_t-1} \frac{1}{(\alpha_s + j)^2} \end{aligned} \quad (22)$$

for $i = 0, 1, \dots, k$ and

$$I_{tii'}(\alpha) = - \sum_{j=0}^{n_t-1} \frac{1}{(\alpha_s + j)^2} \quad (23)$$

for $i, i' = 0, 1, \dots, k$ with $i \neq i'$. Sometimes, $|\mathcal{X}_t|$ is very large at time t in a manufacturing process, e.g., $|\mathcal{X}_t| = 82,408,626,300$ when $n_t = 200$ and $k = 6$. In such situations, it will take too much time to evaluate $I_{tii}(\alpha)$ s by equation (18). One possible approach to evaluate $I_{tii}(\alpha)$ s is the Monte Carlo method as follows: First generate *i.i.d.* $(p_t^{(1)T}, x_t^{(1)T})^T, \dots, (p_t^{(r)T}, x_t^{(r)T})^T$ such that $p_t^{(u)}$ is sampled from $\text{Dirichlet}(\alpha)$ and $x_t^{(u)}$ given $p_t^{(u)}$ is sampled from $\text{multinomial}(n_t; p_t^{(u)})$ for $u = 1, \dots, r$, where r is a large positive integer, e.g., $r = 100,000$. Then $I_{tii}(\alpha)$ can be approximately evaluated by $\widehat{I_{tii}(\alpha)}$, where

$$\widehat{I_{tii}(\alpha)} \equiv \frac{1}{r} \sum_{u=1}^r \left(\sum_{j=0}^{x_{it}^{(u)}-1} \frac{1}{\alpha_i + j} - \sum_{j=0}^{n_t-1} \frac{1}{\alpha_s + j} \right)^2 \quad (24)$$

or

$$\frac{1}{r} \sum_{u=1}^r \sum_{j=0}^{x_{it}^{(u)}-1} \frac{1}{(\alpha_i + j)^2} - \sum_{j=0}^{n_t-1} \frac{1}{(\alpha_s + j)^2}. \quad (25)$$

3 Multivariate Exponentially Weighted Moving Average Monitoring Scheme

In this section, a multivariate exponentially weighted moving average control chart is proposed for monitoring a categorical manufacturing process under the Dirichlet-compound multinomial model as follows: Suppose that $\{(p_t^T, x_t^T)^T: t \geq 1\}$ are independent $(2k+2) \times 1$ random vectors, where both p_t and x_t are described in Section 2. Let $\alpha_0 \in (0, \infty)^{k+1}$ denote the known $(k+1) \times 1$ in-control process parameter vector in phase I. The multivariate exponentially weighted moving average (MEWMA) control chart is based on the $(k+1) \times 1$ vectors

$$\begin{aligned} w_t &\equiv (I_{k+1} - R)w_{t-1} + RS(\alpha_0; x_t) \\ &= R(I_{k+1} - R)^{t-1}S(\alpha_0; x_1) + \dots + R(I_{k+1} - R)S(\alpha_0; x_{t-1}) + RS(\alpha_0; x_t) \\ &= R \sum_{i=0}^{t-1} (I_{k+1} - R)^i S(\alpha_0; x_{t-i}) \end{aligned} \quad (26)$$

for $t = 1, 2, \dots$, where $w_0 \equiv 0_{(k+1) \times 1}$; I_{k+1} denotes the identity matrix of order $k+1$; R is a specified $(k+1) \times (k+1)$ positive definite covariance matrix such that $I_{k+1} - R$ is non-negative definite; and $(I_{k+1} - R)^0 \equiv I_{k+1}$. It follows from the eigenvalue decomposition that $R = P\Lambda P^T$ where P is an orthogonal matrix, i.e., $PP^T = I_{k+1}$, and Λ is a diagonal matrix $\text{diag}\{\lambda_0, \lambda_1, \dots, \lambda_k\}$ with $\lambda_0, \lambda_1, \dots, \lambda_k \in (0, 1]$. Then

$$w_t = P\Lambda \sum_{i=0}^{t-1} (I_{k+1} - \Lambda)^i P^T S(\alpha_0; x_{t-i}) \quad (27)$$

for $t = 1, 2, \dots$

When $p_1, \dots, p_t \sim \text{Dirichlet}(\alpha_0)$ for some $t = 1, 2, \dots$, the following properties hold:

(i) $E_{\alpha_0}(w_t) = 0_{(k+1) \times 1}$.

(ii)

$$\text{Cov}_{\alpha_0}(w_t) = P\Lambda \left[\sum_{i=0}^{t-1} (I_{k+1} - \Lambda)^i P^T I_{t-i}(\alpha_0) P (I_{k+1} - \Lambda)^i \right] \Lambda P^T (\equiv \Sigma_t). \quad (28)$$

(iii) $E_{\alpha_0}(w_t^T \Sigma_t^{-1} w_t) = k+1$. Set

$$T_t^2 \equiv w_t^T \Sigma_t^{-1} w_t. \quad (29)$$

(iv)

$$(R^{-1}\Sigma_t R^{-1})^{-1/2} R^{-1} w_t \xrightarrow{a.s.} \left[\sum_{i=0}^{t-1} I_{t-i}(\alpha_0) \right]^{-1/2} \sum_{i=0}^{t-1} S(\alpha_0; x_{t-i}) \quad (30)$$

and

$$T_t^2 \xrightarrow{a.s.} \left[\sum_{i=0}^{t-1} S(\alpha_0; x_{t-i}) \right]^T \left[\sum_{i=0}^{t-1} I_{t-i}(\alpha_0) \right]^{-1} \sum_{i=0}^{t-1} S(\alpha_0; x_{t-i}) (\equiv T_{t0}^2) \quad (31)$$

as $\max\{\lambda_0, \lambda_1, \dots, \lambda_k\} \rightarrow 0$, where T_{t0}^2 is the score test statistic up to time t for testing the null hypothesis $H_0: p_1, \dots, p_t \sim \text{Dirichlet}(\alpha_0)$ versus the alternative $H_1: p_1, \dots, p_t \sim \text{Dirichlet}(\alpha)$ for some $\alpha \neq \alpha_0$.

(v) If $\sup_{t \geq 1} n_t < \infty$, then $(R^{-1}\Sigma_t R^{-1})^{-1/2} R^{-1} w_t \xrightarrow{d} N_{k+1}(0_{(k+1) \times 1}, I_{k+1})$ and $T_t^2 \xrightarrow{d} \chi_{k+1}^2$ as $\max\{\lambda_0, \lambda_1, \dots, \lambda_k\} \rightarrow 0$ and $t \rightarrow \infty$.

(vi) If $\Lambda = \lambda I_{k+1}$ for some $\lambda \in (0, 1]$, then

$$w_t = \lambda \sum_{i=0}^{t-1} (1-\lambda)^i S(\alpha_0; x_{t-i}), \quad (32)$$

$$T_t^2 = \left[\sum_{i=0}^{t-1} (1-\lambda)^i S(\alpha_0; x_{t-i}) \right]^T \left[\sum_{i=0}^{t-1} (1-\lambda)^{2i} I_{t-i}(\alpha_0) \right]^{-1} \sum_{i=0}^{t-1} (1-\lambda)^i S(\alpha_0; x_{t-i}), \quad (33)$$

and

$$\Sigma_t = \lambda^2 \sum_{i=0}^{t-1} (1-\lambda)^{2i} I_{t-i}(\alpha_0) \quad (34)$$

for $t = 1, 2, \dots$, where $0^0 \equiv 1$. In particular, if $\Lambda = I_{k+1}$, then $w_t = S(\alpha_0; x_t)$, $T_t^2 = S^T(\alpha_0; x_t) I_t^{-1}(\alpha_0) S(\alpha_0; x_t) (\equiv T_{t1}^2)$, and $\Sigma_t = I_t(\alpha_0)$ for $t = 1, 2, \dots$, where T_{t1}^2 is the score test statistic at time t for testing the null hypothesis $H_0: p_t \sim \text{Dirichlet}(\alpha_0)$ versus the alternative $H_1: p_t \sim \text{Dirichlet}(\alpha)$ for some $\alpha \neq \alpha_0$.

(vii) If $n_1 = n_2 = \dots$, then

$$\begin{aligned} \Sigma_t &= P \Lambda \left[\sum_{i=0}^{t-1} (I_{k+1} - \Lambda)^i P^T I_1(\alpha_0) P (I_{k+1} - \Lambda)^i \right] \Lambda P^T \\ &\equiv P \Lambda \left[\sum_{i=0}^{t-1} (I_{k+1} - \Lambda)^i I_1^*(\alpha_0) (I_{k+1} - \Lambda)^i \right] \Lambda P^T \\ &= P \left(\frac{\lambda_j \lambda_{j'} [1 - (1-\lambda_j)^t (1-\lambda_{j'})^t] I_{1j j'}^*(\alpha_0)}{\lambda_j + \lambda_{j'} - \lambda_j \lambda_{j'}} \right)_{j, j'=0, 1, \dots, k} P^T \\ &\rightarrow P \left(\frac{\lambda_j \lambda_{j'} I_{1j j'}^*(\alpha_0)}{\lambda_j + \lambda_{j'} - \lambda_j \lambda_{j'}} \right)_{j, j'=0, 1, \dots, k} P^T \end{aligned} \quad (35)$$

as $t \rightarrow \infty$, where $P^T I_1(\alpha_0) P \equiv I_1^*(\alpha_0) \equiv (I_{1jj'}^*(\alpha_0))_{j,j'=0,1,\dots,k}$.

(viii) If $n_1 = n_2 = \dots$ and $\Lambda = \lambda I_{k+1}$ for some $\lambda \in (0, 1]$, then

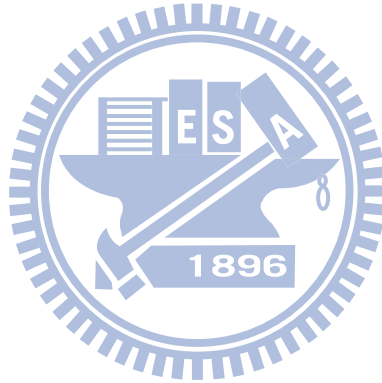
$$\Sigma_t = \frac{\lambda[1 - (1 - \lambda)^{2t}]}{2 - \lambda} I_1(\alpha_0) \rightarrow \frac{\lambda}{2 - \lambda} I_1(\alpha_0) \quad (36)$$

as $t \rightarrow \infty$. In particular, if $n_1 = n_2 = \dots$ and $\Lambda = I_{k+1}$, then $\Sigma_t = I_1(\alpha_0)$ for $t = 1, 2, \dots$, w_1, w_2, \dots are *i.i.d.*, and T_1^2, T_2^2, \dots are *i.i.d.*

Then the stopping time of the MEWMA monitoring scheme is the first time t such that

$$T_t^2 > h, \quad (37)$$

where h is chosen to achieve a specified in-control average run length, e.g., $1/[2\Phi(-3)] \approx 370.4$ with $\Phi(\cdot)$ denoting the cumulative distribution function (c.d.f.) of the standard normal distribution.



4 A Simulation Study

In the former information, it accumulate the data up to time t when $t = 1, 2, \dots, l$. Through the former information, this paper discusses the different l for each λ . The l is the past in-control data in phase I.

In order to study the performance of this quality control scheme, it compute the average run length. To evaluate the in-control average run length ($\equiv ARL_0$), it considers the special case where $k = 2$; p_1, p_2, \dots are sampled from Dirichlet(α_0) with $\alpha_0 = (85, 10, 5)^T$; $n_1 = n_2 = \dots = 100$; $P = I_3$; and $\Lambda = \lambda I_3$ for $\lambda \in \{0.01, 0.05, 0.10, 0.15, 0.20, 0.30, 0.40, 0.50, 1\}$ ($\equiv S$).

For $\lambda \in S$ and $ARL_0 = 1/[2\Phi(-3)]$, The h in equation (33) can be evaluated as follows:

Step 1: Generate *i.i.d.* $(p_1^T, x_1^T)^T, \dots, (p_{370}^T, x_{370}^T)^T$ such that p_t is sampled from Dirichlet(α_0) and x_t is sampled from multinomial($100; p_t$) for $t = 1, 2, \dots, 370$.

Step 2: To sort the T_t^2 s such that $T_t^{2(1)}, T_t^{2(2)}, \dots, T_t^{2(370)}$.

Step 3: Choose the maximum $T_t^{2(370)}$ ($\equiv T_t^{*2}$).

Repeat Steps 1–3 for 10,000 times independently. To sort the T_t^{*2} s such that $T_t^{*2(1)}, T_t^{*2(2)}, \dots, T_t^{*2(10,000)}$. The initial value of h_1 is $T_t^{*2(5,000)}$. To compute the $ARL_0^{(1)}$ with h_1 . If the $ARL_0^{(1)}$ is large than $1/[2\Phi(-3)]$, then h_2 is given such that $h_2 < h_1$, and compute the $ARL_0^{(2)}$. If the $ARL_0^{(1)}$ is smaller than $1/[2\Phi(-3)]$, then h_2 is given such that $h_2 > h_1$, and compute the $ARL_0^{(2)}$. Until it finds the h such that $ARL_0 \approx 1/[2\Phi(-3)]$ for each λ ; see Table 1.

To evaluate the out-of-control average run length ($\equiv ARL_1$), it considers the special case where $k = 2$; p_1, \dots, p_l are sampled from Dirichlet(α_0) and p_{l+1}, p_{l+2}, \dots are sampled from Dirichlet(α_1) for some $l \in \{0, 1, 2, \dots, 500\}$ with $\alpha_0 = (85, 10, 5)^T$, $\alpha_1 \in \{(80, 12.5, 7.5)^T (\equiv \alpha_1^{(1)}), (75, 15, 10)^T (\equiv \alpha_1^{(2)}), (70, 20, 10)^T (\equiv \alpha_1^{(3)})\}$; $n_1 = n_2 = \dots = 100$; $P = I_3$; and $\Lambda = \lambda I_3$ for $\lambda \in S$.

The ARL_1 can be evaluated as follows:

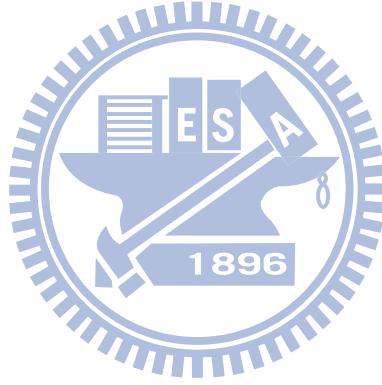
Step 1: Generate *i.i.d.* $(p_1^T, x_1^T)^T, \dots, (p_l^T, x_l^T)^T$ such that p_t is sampled from Dirichlet(α_0) and then x_t is sampled from multinomial($100; p_t$) for $t = 1, 2, \dots, l$.

Step 2: If $T_t^2 > h$ for some $t \in \{1, 2, \dots, l\}$, then return to Step 1.

Step 3: Generate *i.i.d.* $(p_{l+1}^T, x_{l+1}^T)^T, \dots, (p_{l+t^*}^T, x_{l+t^*}^T)^T$ such that p_t is sampled from $\text{Dirichlet}(\alpha_1)$ and then x_t is sampled from $\text{multinomial}(100; p_t)$ for $t = l+1, l+2, \dots, l+t^*$, where $l+t^*$ is the first time t such that $T_t^2 > h$.

Repeat Steps 1–3 for 100,000 times independently. Then the ARL_1 is approximated evaluated by the average of 100,000 t^* s for each $\lambda \in S$.

For $\lambda \in S$ and $ARL_0 = 1/[2\Phi(-3)]$, the ARL_1 is put in Tables 2–31. The proposed MEWMA monitoring scheme is also compared with the following monitoring scheme: the stopping time of the monitoring scheme is the first time t such that $T_{t_0}^2 > h_0$, where $T_{t_0}^2$ is defined in equation (27) and h_0 is chosen to achieve $ARL_0 = 1/[2\Phi(-3)]$.



5 Comparison and Conclusions

In the paper, an MEWMA control chart is proposed for monitoring a manufacturing process in the Dirichlet-compound multinomial model. For each $\lambda \in S \cup \{0\}$ ($\equiv S'$), it is seen that h increases as λ increases; see Table 1.

For $0 \leq l \leq 4$ and $\lambda \in S'$, the ARL_1 increases as λ increases; see Tables 2–6. Then λ is chosen as 0.

For $l \geq 5$ and $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$, consider the following three cases:

Case 1: $\alpha_1 = \alpha_1^{(1)}$. When $5 \leq l \leq 20$, the MEWMA monitoring scheme with $\lambda = 0$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 7–22. When $21 \leq l \leq 36$, the MEWMA monitoring scheme with $\lambda = 0.05$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 23–24. When $38 \leq l \leq 500$, the MEWMA monitoring scheme with $\lambda = 0.1$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 26–31.

Case 2: $\alpha_1 = \alpha_1^{(2)}$. When $5 \leq l \leq 11$, the MEWMA monitoring scheme with $\lambda = 0$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 7–13. When $12 \leq l \leq 138$, the MEWMA monitoring scheme with $\lambda = 0.15$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 14–27. When $140 \leq l \leq 500$, the MEWMA monitoring scheme with $\lambda = 0.2$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 29–31.

Case 3: $\alpha_1 = \alpha_1^{(3)}$. When $5 \leq l \leq 8$, the MEWMA monitoring scheme with $\lambda = 0$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 7–10. But when the $9 \leq l \leq 500$, the MEWMA monitoring scheme with $\lambda = 0.3$ has the smallest ARL_1 for $\lambda \in S'$; see Tables 11–31.

For these three cases, it is seen that when α_1 is faraway from α_0 , e.g., $\alpha_1 = \alpha_1^{(3)}$, the ARL_1 is smaller. When α_1 is close to α_0 , e.g., $\alpha_1 = \alpha_1^{(1)}$, the ARL_1 is bigger.

Compare $\lambda = 0.05$ and 0.01:

Case 1: $\alpha_1 = \alpha_1^{(1)}$. When $0 \leq l \leq 15$, the ARL_1 for $\lambda = 0.01$ is smaller than that for $\lambda = 0.05$, so $\lambda = 0.01$ is better than $\lambda = 0.05$; see Tables 2–17. When $16 \leq l \leq 500$, the ARL_1 for $\lambda = 0.05$ is smaller than that for $\lambda = 0.01$, so $\lambda = 0.05$ is better than $\lambda = 0.01$; see Tables 18–31.

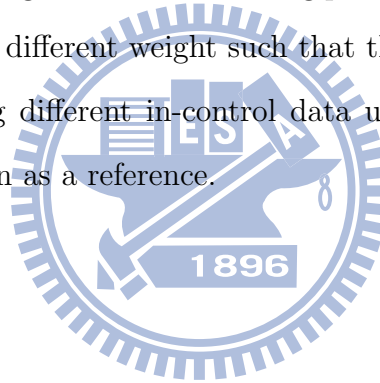
Case 2: $\alpha_1 = \alpha_1^{(2)}$. When $0 \leq l \leq 12$, the ARL_1 for $\lambda = 0.01$ is smaller than that for $\lambda = 0.05$, so $\lambda = 0.01$ is better than $\lambda = 0.05$; see Tables 2–14. When $13 \leq l \leq 500$, the

ARL_1 for $\lambda = 0.05$ is smaller than that for $\lambda = 0.01$, so $\lambda = 0.05$ is better than $\lambda = 0.01$; see Tables 14–31.

Case 3: $\alpha_1 = \alpha_1^{(3)}$. When $0 \leq l \leq 11$, the ARL_1 for $\lambda = 0.01$ is smaller than that for $\lambda = 0.05$, so $\lambda = 0.01$ is better than $\lambda = 0.05$; see Tables 2–13. When $12 \leq l \leq 500$, the ARL_1 for $\lambda = 0.05$ is smaller than that for $\lambda = 0.01$, so $\lambda = 0.05$ is better than $\lambda = 0.01$; see Tables 14–31.

So, when l is large, the weight $\lambda = 0.05$ is better than $\lambda = 0.01$.

In general, the λ of EWMA in the interval $0.05 \leq \lambda \leq 0.25$ works well in practice, with $\lambda = 0.05$, $\lambda = 0.10$, $\lambda = 0.20$ being popular choices. In this paper, it suggests a multivariate exponentially weighted moving average control chart for some relevant properties of the proposed control chart are also investigated in the Dirichlet-compound multinomial model for a categorical manufacturing process. It is found by simulation that the different α_1 and l have different weight such that the ARL_1 is smallest when $\lambda \in S'$. According to accumulating different in-control data up to time l in phase I, λ will be different. This can be taken as a reference.



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Table 1: h for $\lambda \in S'$ and $\alpha_0 = (85, 10, 5)^T$.

λ	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
h	6.53	8.33	11.96	14.79	17.02	19.08	22.45	25.42	27.73	34.34

* denotes the smallest ARL_1 .

Table 2: ARL_1 for $l = 0$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$, where $\alpha_1^{(1)} \equiv (80, 12.5, 7.5)^T$, $\alpha_1^{(2)} \equiv (75, 15, 10)^T$, and $\alpha_1^{(3)} \equiv (70, 20, 10)^T$.

$l = 0$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	5.00*	6.19	8.32	10.10	12.00	14.08	18.64	23.47	27.90	45.20
$(75, 15, 10)^T$	1.94*	2.19	2.62	2.96	3.22	3.49	3.96	4.52	5.13	8.80
$(70, 20, 10)^T$	1.30*	1.39	1.54	1.66	1.77	1.86	1.99	2.16	2.28	3.32

Table 3: ARL_1 for $l = 1$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 1$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	6.00*	7.20	9.29	10.94	12.71	14.76	19.13	23.75	28.03	45.2
$(75, 15, 10)^T$	2.37*	2.65	3.14	3.44	3.65	3.86	4.26	4.75	5.29	8.80
$(70, 20, 10)^T$	1.53*	1.64	1.84	1.96	2.05	2.11	2.21	2.31	2.41	3.32

Table 4: ARL_1 for $l = 2$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 2$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	6.64*	7.94	9.83	11.31	13.06	14.94	19.23	24.10	27.87	45.20
$(75, 15, 10)^T$	2.64*	2.96	3.44	3.69	3.86	4.04	4.38	4.82	5.03	8.80
$(70, 20, 10)^T$	1.69*	1.82	2.03	2.14	2.19	2.24	2.29	2.37	2.44	3.32

Table 5: ARL_1 for $l = 3$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 3$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	7.17*	8.36	10.19	11.69	13.24	15.05	19.38	24.26	28.03	45.20
$(75, 15, 10)^T$	2.89*	3.22	3.66	3.87	3.99	4.11	4.41	4.84	5.31	8.80
$(70, 20, 10)^T$	1.82*	1.97	2.17	2.25	2.28	2.30	2.33	2.37	2.44	3.32

Table 6: ARL_1 for $l = 4$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 4$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	7.56*	8.90	10.52	11.76	13.24	15.27	19.52	24.07	27.93	45.20
$(75, 15, 10)^T$	3.09*	3.42	3.84	4.01	4.08	4.18	4.43	4.85	5.32	8.80
$(70, 20, 10)^T$	1.94*	2.10	2.29	2.34	2.34	2.35	2.35	2.38	2.44	3.32

Table 7: ARL_1 for $l = 5$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 5$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	8.01*	9.24	10.85	11.94	13.58	15.18	19.57	24.12	28.19	45.20
$(75, 15, 10)^T$	3.27*	3.63	3.99	4.10	4.13	4.23	4.46	4.87	5.37	8.80
$(70, 20, 10)^T$	2.05*	2.21	2.38	2.41	2.40	2.37	2.35	2.38	2.45	3.32

Table 8: ARL_1 for $l = 6$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 6$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	8.40*	9.55	11.14	12.05	13.49	15.21	19.49	24.14	28.09	45.20
$(75, 15, 10)^T$	3.45*	3.78	4.12	4.16	4.18	4.22	4.44	4.85	5.34	8.80
$(70, 20, 10)^T$	2.16*	2.31	2.46	2.46	2.42	2.39	2.37	2.38	2.44	3.32

Table 9: ARL_1 for $l = 7$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 7$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	8.80*	9.95	11.23	12.07	13.46	15.33	19.37	24.16	28.44	45.20
$(75, 15, 10)^T$	3.58*	3.93	4.23	4.22	4.20	4.24	4.44	4.85	5.36	8.80
$(70, 20, 10)^T$	2.23*	2.40	2.53	2.50	2.43	2.40	2.37	2.38	2.44	3.32

Table 10: ARL_1 for $l = 8$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 8$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	9.10*	10.10	11.31	12.14	13.43	15.30	19.17	24.17	27.59	45.20
$(75, 15, 10)^T$	3.75*	4.09	4.33	4.27	4.22	4.27	4.44	4.85	5.37	8.80
$(70, 20, 10)^T$	2.33*	2.49	2.60	2.53	2.46	2.40	2.36	2.38	2.44	3.32

Table 11: ARL_1 for $l = 9$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 9$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	9.40*	10.53	11.51	12.39	13.58	15.28	19.64	24.01	27.79	45.20
$(75, 15, 10)^T$	3.91*	4.21	4.41	4.29	4.22	4.23	4.45	4.84	5.36	8.80
$(70, 20, 10)^T$	2.40	2.57	2.65	2.56	2.48	2.40	2.36*	2.38	2.45	3.32

Table 12: ARL_1 for $l = 10$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 10$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	9.68*	10.79	11.62	12.26	13.55	15.35	19.35	23.99	28.01	45.20
$(75, 15, 10)^T$	4.01*	4.35	4.48	4.34	4.23	4.25	4.46	4.86	5.34	8.80
$(70, 20, 10)^T$	2.50	2.64	2.69	2.58	2.47	2.40	2.35*	2.37	2.45	3.32

Table 13: ARL_1 for $l = 11$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 11$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	10.03*	11.03	11.79	12.31	13.57	15.33	19.55	24.17	28.23	45.20
$(75, 15, 10)^T$	4.16*	4.45	4.55	4.36	4.27	4.25	4.46	4.85	5.33	8.80
$(70, 20, 10)^T$	2.55	2.71	2.75	2.59	2.48	2.41	2.36*	2.38	2.44	3.32

Table 14: ARL_1 for $l = 12$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 12$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	10.24*	11.31	11.82	12.35	13.55	15.31	19.60	24.19	28.19	45.20
$(75, 15, 10)^T$	4.27	4.57	4.59	4.38	4.24*	4.26	4.44	4.85	5.35	8.80
$(70, 20, 10)^T$	2.64	2.78	2.77	2.61	2.48	2.41	2.36*	2.39	2.44	3.32

Table 15: ARL_1 for $l = 13$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 13$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	10.53*	11.54	11.93	12.41	13.56	15.29	19.32	24.23	28.40	45.20
$(75, 15, 10)^T$	4.40	4.68	4.65	4.39	4.26*	4.27	4.44	4.85	5.36	8.80
$(70, 20, 10)^T$	2.70	2.85	2.80	2.61	2.47	2.42	2.35*	2.38	2.44	3.32

Table 16: ARL_1 for $l = 14$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 14$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	10.70*	11.63	11.97	12.40	13.56	15.31	19.23	24.15	28.32	45.20
$(75, 15, 10)^T$	4.48	4.77	4.71	4.41	4.27*	4.28	4.43	4.86	5.34	8.80
$(70, 20, 10)^T$	2.76	2.92	2.81	2.62	2.47	2.41	2.37*	2.38	2.44	3.32

Table 17: ARL_1 for $l = 15$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 15$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	10.96*	11.81	12.03	12.39	13.57	15.33	19.37	24.17	28.28	45.20
$(75, 15, 10)^T$	4.59	4.86	4.74	4.42	4.27*	4.28	4.45	4.86	5.35	8.80
$(70, 20, 10)^T$	2.82	2.97	2.84	2.62	2.48	2.41	2.35*	2.38	2.43	3.32

Table 18: ARL_1 for $l = 16$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 16$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	11.22*	12.02	12.00	12.38	13.58	15.29	19.24	24.26	28.17	45.20
$(75, 15, 10)^T$	4.68	4.95	4.77	4.42	4.26*	4.27	4.44	4.87	5.35	8.80
$(70, 20, 10)^T$	2.88	3.05	2.87	2.63	2.47	2.42	2.33*	2.37	2.44	3.32

Table 19: ARL_1 for $l = 17$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 17$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	11.50*	12.37	12.16	12.38	13.56	15.31	19.30	24.03	28.25	45.20
$(75, 15, 10)^T$	4.81	5.07	4.79	4.43	4.26*	4.27	4.44	4.85	5.34	8.80
$(70, 20, 10)^T$	2.95	3.08	2.90	2.64	2.48	2.42	2.34*	2.38	2.44	3.32

Table 20: ARL_1 for $l = 18$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 18$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	11.77*	12.61	12.18	12.38	13.55	15.34	19.17	23.92	28.33	45.20
$(75, 15, 10)^T$	4.91	5.13	4.82	4.43	4.19*	4.26	4.45	4.85	5.35	8.80
$(70, 20, 10)^T$	3.01	3.11	2.91	2.64	2.49	2.41	2.33*	2.38	2.45	3.32

Table 21: ARL_1 for $l = 19$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 19$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	12.05*	12.64	12.22	12.52	13.56	15.32	19.23	24.29	28.39	45.20
$(75, 15, 10)^T$	5.01	5.25	4.82	4.44	4.25*	4.27	4.44	4.86	5.37	8.80
$(70, 20, 10)^T$	3.07	3.15	2.93	2.63	2.48	2.42	2.36*	2.39	2.45	3.32

Table 22: ARL_1 for $l = 20$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 20$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	12.12*	12.84	12.23	12.49	13.58	15.33	19.33	24.03	28.24	45.20
$(75, 15, 10)^T$	5.10	5.26	4.90	4.45	4.25*	4.28	4.44	4.85	5.36	8.80
$(70, 20, 10)^T$	3.09	3.21	2.94	2.63	2.49	2.42	2.37*	2.39	2.44	3.32

Table 23: ARL_1 for $l = 21$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 21$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	12.42	13.01	12.27*	12.43	13.57	15.29	19.20	23.94	28.13	45.20
$(75, 15, 10)^T$	5.17	5.38	4.91	4.44	4.26*	4.27	4.45	4.87	5.35	8.80
$(70, 20, 10)^T$	3.20	3.28	2.95	2.64	2.50	2.43	2.36*	2.37	2.44	3.32

Table 24: ARL_1 for $l = 36$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 36$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	15.63	14.09	12.34*	12.45	13.62	15.35	19.40	24.17	28.18	45.20
$(75, 15, 10)^T$	5.98	6.08	4.95	4.44	4.26*	4.30	4.44	4.86	5.34	8.80
$(70, 20, 10)^T$	3.77	3.69	3.01	2.65	2.50	2.43	2.34*	2.38	2.44	3.32

Table 25: ARL_1 for $l = 37$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 37$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	15.76	14.21	12.46*	12.46*	13.65	15.37	19.43	24.19	28.20	45.20
$(75, 15, 10)^T$	6.11	6.14	4.99	4.44	4.27*	4.29	4.44	4.86	5.36	8.80
$(70, 20, 10)^T$	3.82	3.73	3.03	2.66	2.50	2.43	2.35*	2.38	2.44	3.32

Table 26: ARL_1 for $l = 38$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 38$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	16.15	14.37	12.51	12.47*	13.63	15.39	19.47	24.04	28.17	45.20
$(75, 15, 10)^T$	6.24	6.19	5.00	4.44	4.27*	4.31	4.44	4.87	5.33	8.80
$(70, 20, 10)^T$	3.90	3.76	3.04	2.65	2.50	2.43	2.36*	2.37	2.44	3.32

Table 27: ARL_1 for $l = 138$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 138$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	26.84	18.16	12.63	12.39*	13.60	15.30	19.42	23.94	28.19	45.20
$(75, 15, 10)^T$	11.32	7.78	5.07	4.47	4.28*	4.31	4.44	4.85	5.35	8.80
$(70, 20, 10)^T$	6.71	4.67	3.07	2.65	2.49	2.41	2.35*	2.38	2.45	3.32

Table 28: ARL_1 for $l = 139$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 139$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	26.91	18.17	12.63	12.41*	13.59	15.31	19.43	24.06	28.17	45.20
$(75, 15, 10)^T$	11.36	7.79	5.07	4.46	4.26*	4.26*	4.45	4.86	5.36	8.80
$(70, 20, 10)^T$	6.74	4.68	3.06	2.65	2.50	2.42	2.36*	2.38	2.44	3.32

Table 29: ARL_1 for $l = 140$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 140$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	27.04	18.18	12.64	12.42*	13.58	15.31	19.47	24.07	28.17	45.20
$(75, 15, 10)^T$	11.38	7.79	5.06	4.45	4.32	4.21*	4.45	4.86	5.36	8.80
$(70, 20, 10)^T$	6.75	4.68	3.05	2.66	2.48	2.42	2.37*	2.39	2.44	3.32

Table 30: ARL_1 for $l = 450$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 450$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	45.17	19.08	12.75	12.61*	13.55	15.38	19.45	24.34	28.52	45.20
$(75, 15, 10)^T$	19.36	8.34	5.11	4.54	4.47	4.32*	4.46	5.09	5.39	8.80
$(70, 20, 10)^T$	11.71	4.87	3.10	2.68	2.51	2.42	2.33*	2.41	2.40	3.32

Table 31: ARL_1 for $l = 500$, $\lambda \in S'$; $\alpha_1 \in \{\alpha_1^{(1)}, \alpha_1^{(2)}, \alpha_1^{(3)}\}$.

$l = 500$	λ									
	0.00	0.01	0.05	0.10	0.15	0.20	0.30	0.40	0.50	1.00
$(80, 12.5, 7.5)^T$	48.04	19.30	12.80	12.67*	13.57	15.42	19.44	24.38	28.49	45.20
$(75, 15, 10)^T$	20.21	8.42	5.19	4.53	4.49	4.36*	4.47	5.13	5.42	8.80
$(70, 20, 10)^T$	12.09	4.89	3.11	2.66	2.52	2.40	2.31*	2.43	2.36	3.32

